## 2005 Pollock Year-Class Prediction: Average Recruitment

31 August 2005

## DATA

This forecast is based on five data sources: three physical properties and two biological data sets. The sources are:

1. Observed 2005 Kodiak monthly precipitation. The Kodiak Weather Service Office (http://padq.arh.noaa.gov/) prepares monthly precipitation totals (inches) from hourly observations. Data for 2005 were obtained from the NOAA National Climate Data Center, Asheville, North Carolina.
2. Wind mixing energy at $\left[57^{\circ} \mathrm{N}, 156^{\circ} \mathrm{W}\right]$ estimated from 2005 sea-level pressure analyses. Monthly estimates of wind mixing energy ( $\mathrm{W} \mathrm{m}^{-2}$ ) were computed for a location near the southwestern end of Shelikof Strait. To make the estimates, twice-daily gradient winds were computed for that location using the METLIB utility (Macklin et al., 1984). Gradient winds were converted to surface winds using an empirical formula based on Macklin et al. (1993). Estimates of wind mixing energy were computed using constant air density ( $1.293 \mathrm{~kg} \mathrm{~m}^{-3}$ ) and the drag coefficient formulation of Large and Pond (1982).
3. Advection of ocean water near Shelikof Strait inferred from drogued drifters deployed during the spring of 2005.
4. Rough counts of pollock larvae from a survey conducted in late May-early June 2005.
5. Estimates of age-2 pollock abundance and spawner biomass from the 2005 assessment.


#### Abstract

ANALYSIS Kodiak Precipitation: Kodiak precipitation is a proxy for fresh-water runoff that contributes to the density contrast between coastal and Alaska Coastal Current water in Shelikof Strait. The greater the contrast, the more likely that eddies and other instabilities will form. Such secondary circulations have attributes that make them beneficial to survival of larval pollock.

The season began with typical precipitation during January (Table 1). For all contributing winter and spring months, precipitation was near or above normal, with February being the wettest (at $153 \%$ of the $30-\mathrm{yr}$ February average.


TABLE 1. Kodiak precipitation for 2005.

| Month | \% 30-yr average |
| :---: | :---: |
| Jan | 104 |
| Feb | 153 |
| Mar | 111 |
| Apr | 103 |
| May | 139 |
| June | 104 |

Based on this information, the forecast element for Kodiak 2005 rainfall has a score of 2.21. This is "average to strong" on the continuum from 1 (weak) to 3 (strong).

Wind Mixing: Following the decadal trend established in the late 1990s, wind mixing at the southern end of Shelikof Strait was again below the long-term average for all winter and spring months of 2005, except March.

TABLE 2. Wind mixing at the exit of Shelikof Strait for 2005.

| Month | \% 30-yr average |
| :---: | :---: |
| Jan | 46 |
| Feb | 48 |
| Mar | 114 |
| Apr | 74 |
| May | 39 |
| June | 39 |

Strong mixing in winter helps transport nutrients into the upper ocean layer to provide a basis for the spring phytoplankton bloom. Weak spring mixing is thought to better enable first feeding pollock larvae to locate and capture food. Weak mixing in winter is not conducive to high survival rates, while weak mixing in spring favors recruitment. This year's scenario produces a wind mixing score of 2.29 , which equates to "average-to-strong".

Advection: From an examination of drifter trajectories and wind forcing, the transport in Shelikof Strait for spring of 2005 was strong until mid April and then weak, which would support a prediction of an average to strong year class.

We have hypothesized that very strong transport is bad for pollock survival, that moderate transport is best, and that very weak transport, while not as disastrous as strong transport, still is detrimental to larval survival. Advection was given a score of 2.29.

Relating the Larval Index to Recruitment: As in last year's analysis, a nonlinear neural network model with one input neuron (larval abundance), three hidden neurons, and one output neuron (recruitment) was used to relate larval abundance (CPUA, average catch, $\mathrm{m}^{-2}$ ) to age-2 recruitment abundance (billions). The model estimated six weighting parameters.

The neural network model, which used the 19 observation pairs of Table 3 to fit the model, had a very low $\mathrm{R}^{2}$ of 0.054 . A plot of the observed recruitment (actual) and that predicted from larval abundance (predicted) are given in Fig. 1, where row number corresponds to the rows of the data matrix given in Table 3.

TABLE 3. Data used in the neural network model.

| Year Class | Mean <br> CPUA | Recruit |
| :---: | :---: | :---: |
| 1982 | 71.14483 | 0.206506 |
| 1985 | 80.42379 | 0.539391 |
| 1987 | 329.7428 | 0.361222 |
| 1988 | 217.9464 | 1.60372 |
| 1989 | 537.2899 | 1.04255 |
| 1990 | 373.8137 | 0.418636 |
| 1991 | 54.21859 | 0.239326 |
| 1992 | 562.7872 | 0.141279 |
| 1993 | 185.3388 | 0.212236 |
| 1994 | 126.5823 | 0.828361 |
| 1995 | 605.2316 | 0.402497 |
| 1996 | 477.6918 | 0.172455 |
| 1997 | 568.421 | 0.179436 |
| 1998 | 74.29526 | 0.266972 |
| 1999 | 119.071 | 1.17074 |
| 2000 | 492.0364 | 0.734729 |
| 2001 | 171.3022 | 0.103318 |
| 2002 | 175.6366 | 0.074741 |
| 2003 | 133.4611 | 0.188679 |



FIGURE 1. Observed and predicted recruitment values from the larval index-recruitment neural network model.

The trained network was then used to predict the recruitment for 2004 and 2005. The predictions are given in Table 4.

TABLE 4. Neural network model predictions for 2004 and 2005.

| Year | Actual <br> Recruitment | Predicted <br> Recruitment |
| :---: | :---: | :---: |
| 2004 | $\mathrm{n} / \mathrm{a}$ | 0.248256 |
| 2005 | $\mathrm{n} / \mathrm{a}$ | 0.339102 |

These values, using the $33 \%$ ( 0.335393 ) and $66 \%(0.70132)$ cutoff points given below, correspond to a weak 2004 year class and an average 2005 year class.

Larval Index Counts: Plotting the data by year and binning the data into catch $/ 10 \mathrm{~m}^{2}$ categories (given below) provides another view of the data. The pattern for 2005 (based on rough counts) show patterns similar to last year in that most of the data fall into the three lowest binning categories, but there were some data observation occupying the higher density bins. These patterns indicate that the 2005-year class may be below average.


FIGURE 2. A series of histograms for larval walleye pollock densities in late May from 1982 to 2005. Data were binned into catch $10 \mathrm{~m}^{2}$ categories. The data from 2000-2005 are rough counts taken at sea, and the 2005 data are from the 6MF05 cruise that was completed on June 3.

The data for Figure 3 and 4 are taken from a reference area that is routinely sampled and that usually contains the majority of the larvae. This year's distribution of pollock (Fig. 4) appears to be centered in the typical reference area, and the larval abundance figures in the middle of the reference area seem to be average. Also, the distribution of larvae in 2005 (Fig. 4) are further to the west compared to 2004 (Fig.3) suggesting that some of the Shelikof larvae might be in their nursery area at the time of the survey. Comparing the two maps shows that the 2005 rough counts seem to be higher compared to 2004 . Given these two pieces of information, the score for larval index is set to average or 2.0.


FIGURE 3. Mean catch per $10 \mathbf{m}^{2}$ for late May cruises during 1982-2004


FIGURE 4. Mean catch per $10 \mathbf{m}^{2}$ for late May cruises during 1982-2005.

Recruitment Time Series: The time series of recruitment from this year's assessment was analyzed in the context of a probabilistic transition in time. The data set consisted of age 2 abundance estimates from 1961-2005, representing the 1959-2003 year classes. There were a total of 45 recruitment data points. The $33 \%$ ( 0.335393 billion) and $66 \%$ ( 0.70132 billion) percentile cutoff points were calculated from the full time series and used to define the three recruitment states of weak, average and strong. The lower third of the data points were called weak, the middle third average and the upper third strong. Using these definitions, nine transition probabilities were then calculated:

1. Probability of a weak year class following a weak
2. Probability of a weak year class following an average
3. Probability of a weak year class following a strong
4. Probability of an average year class following a weak
5. Probability of an average year class following an average
6. Probability of an average year class following a strong
7. Probability of a strong year class following a weak
8. Probability of a strong year class following an average
9. Probability of a strong year class following a strong

The probabilities were calculated with a time lag of two years so that the 2005 year class could be predicted from the size of the 2003 year class. The 2003 year class was estimated to be 0.188679 billion and was classified as weak. The probabilities of other recruitment states following a weak year class for a lag of 2 years $(\mathrm{n}=45)$ are given below:

TABLE 5. Probability of the 2005 year class being weak, average and strong following a weak 2003 year class.

| 2005 Year Class |  | 2003 Year Class | Probability | $\mathbf{N}$ |
| :---: | :---: | :---: | :---: | :---: |
| Weak | follows | Weak | 0.093 | 4 |
| Average | follows | Weak | 0.070 | 3 |
| Strong | follows | Weak | 0.139 | 6 |

The probability of a strong year class following a weak year class two years later had the highest probability. We classified this data element as a strong, giving it a score at the low end of strong 2.34 .

Spawner/Recruit Time Series: The data from the previous analysis only looked at the time sequence of the recruitment data points. This section looks at both the recruitment (R) and the spawning biomass (SB) in the context of transition probabilities after Rothschild and Mullin (1985). The benefit is that it is non-parametric and it provides a way to predict recruitment without applying a presumed functional spawner-recruit relationship. It involves partitioning the spawning stock into N -tiles and the recruitment into N -tiles, classifying the stock into NxN states. I used the $50 \%$ percentile of the data to calculate the median spawning biomass ( 0.269 million tons) and recruitment ( 0.435 billion). These values were used to partition the spawnerrecruit space into 4 tiles, state 1:low SB-low R, state 2:low SB-high R, state 3:high SB-low R, and state 4:high SB-high R. The classification then makes it possible to study the probability of any state and the transitions between the states.

The time series of recruitment data and the $2 \times 2$ spawning biomass-recruitment plot are shown in Figure 5.


FIGURE 5. Time series of recruitment and the $2 \times 2$ classification of the spawning biomass and recruitment data.

TABLE 6. Transition matrix calculated from data in Figure 5.

| Transition Probability matrix | To state1 | To state 2 | To state 3 | To state 4 |
| :---: | :---: | :---: | :---: | :---: |
| From state 1 | 0.692 | 0.308 | 0.000 | 0.000 |
| From state 2 | 0.375 | 0.500 | 0.000 | 0.125 |
| From state 3 | 0.125 | 0.000 | 0.500 | 0.375 |
| From state 4 | 0.000 | 0.000 | 0.267 | 0.733 |

To calculate the score from Figure 5 takes two steps. First, we determine which state is the current state by taking the estimate of spawning biomass in 2005 ( 0.1827 million tons) and note that it falls below the median value of 0.269 . We can see that in 2005 we are in either state 1 or state 2 . The probabilities of transitioning from state 1 or state 2 to other states are given in the first two rows of Table 6.

If we are in state 1 , then recruitment can either be below (a recruitment score of 1) or above the median (a recruitment score of 3). Note the probability for transitioning from state 1 to state 3 or 4 is 0.0 and from state 2 to state 3 is 0.0 . If we start in state 1 , then the combined recruitment score would be the weighted average of the recruitment scores for each possible transition, where the weighting factors are the probabilities. So, the calculations for the second step proceed as described below.

The weighted recruitment score (given we start in state 1) is the recruitment score for staying in state 1 (recruitment below the median, score $=1$ ) times the weight (the probability of transitioning from state 1 back to state 1) plus the recruitment score for transitioning from state 1 to state 2 (recruitment above the median, score $=3$ ) times the weight (the probability of transitioning from state 1 to state 2 ), all divided by the sum of the weights.

$$
=\frac{(1 * 0.692)+(3 * 0.308)}{(0.692+0.308)}=1.61
$$

Similarly, the weighted recruitment score (given we start in state 2)

$$
=\frac{(1 * 0.375)+(3 * 0.5)+(3 * 0.125)}{(0.375+0.5+0.125)}=2.25
$$

We average over these two weighted scores because stating from either state 1 or state 2 is equally likely if the starting spawning biomass in 2005 is below the median, giving a final score of 1.97, or the middle range of average.

## CONCLUSION

A low weighting score of 0.1 was assigned to the larval index data element because the recruitment variability explained by larval abundance was very low. Each of the remaining data elements were weighted equally.

Based on these six elements and the weights assigned in Table 7, below, the FOCI forecast of the 2005 year class is average.

TABLE 7. Final 2005 pollock recruitment forecast.

| Element | Weights | Score | Total |
| :---: | :---: | :---: | :---: |
| Time Sequence of R | 0.18 | 2.34 | 0.4212 |
| Rain | 0.18 | 2.21 | 0.3978 |
| Wind Mixing | 0.18 | 2.29 | 0.4122 |
| Advection | 0.18 | 2.29 | 0.4122 |
| Larval Index-abundance | 0.10 | 2.00 | 0.2000 |
| Spawner-Recruit Data | 0.18 | 1.68 | 0.3024 |
| Total | $\mathbf{1 . 0 0}$ |  | $\mathbf{2 . 1 4 5 8}=$ <br> Average |

## REFERENCES

Large, W.G., and S. Pond (1982) Sensible and latent heat flux measurement over the ocean. $J$. Phys. Oceanogr. 2: 464-482.

Macklin, S.A., R.L. Brown, J. Gray, and R.W. Lindsay (1984) METLIB-II - A program library for calculating and plotting atmospheric and oceanic fields. NOAA Tech. Memo. ERL PMEL-54, NTIS PB84-205434, 53 pp.

Macklin, S.A., P.J. Stabeno, and J.D. Schumacher (1993) A comparison of gradient and observed over-the-water winds along a mountainous coast. J. Geophys. Res. 98: 16,55516,569.

Rothschild, B. J. and Mullin, A.J. 1985. The information content of stock-and-recruitment data and its non-parametric classification. Journal du Conseil International pour l'Exploration de la Mer. 42: 116-124.

